Estimating abundance, fishing mortality and migration rates by area, using the spatial VPA methodology. Application to yellowfin tuna in eastern and western Atlantic

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INTRODUCTION

Since the agreement of a single Atlantic yellowfin stock (Anon., 1993), current yellowfin stock assessment and management advice are based on population dynamic modelling at the global stock distribution scale. In this way, those methods do not allow spatially heterogeneous assessments. Nevertheless, yellowfin tuna migrate massively at different scales (Bard et Hervé 1994 ; Foucher 1995 ; Fonteneau 1994a ; Fonteneau et al. 1997). Many studies underline the importance of spatial heterogeneity and fish movement for both stock assessment (**) and fleet interaction studies (**). Spatial models do exist and could profitably be applied to Atlantic yellowfin tuna stock. Such models have already been used for different tuna stocks (Porch 1996 ; Sibert et Fournier 1994 ; Kleiber et Fonteneau 1994 ; Butterworth **, ...). But most of them deal with stocks with available tagging data. Unfortunately, because of such data scarceness for Atlantic yellowfin stock, fish migration rates are unknown and spatial stock assessment and modelling remain inapplicable in practice.

This paper uses the spatial VPA to assess fishing mortality rates, fish numbers and migration rates between ICCAT east and west Atlantic zones. Spatial VPA is a method to estimate fish abundance and local fishing mortality coefficients in spatial zones and migration rates between these zones (Maury et al 1997). The method only needs data of catches at age and fishing effort by zone and it takes into account the surface of the stock in each zone. Stock surface is estimated here by long liners catches spatial interpolation using a GIS (Geographical Information System).
1. DATA AND MÉTHOD

1.1. The spatial VPA methodology

The method used here is described in detail in Maury et al. (1997a). The basic model used is a discrete model. It is analogous to the boxes model whose differential formulation was presented by Beverton and Holt in 1957, and is based on the assumption of an instantaneous migration of fishes at the end of each time step (Punt et Butterworth, 1994). Here, the method is applied for two distinct boxes but could be easily extended to more boxes.

\[
\begin{align*}
N_{1,t+1} &= N_{1,t}e^{-\left(F_{1,t}+M_{1,t}\right)}(1-T_{12,t}) + N_{2,t}e^{-\left(F_{2,t}+M_{2,t}\right)}T_{21,t} \\
N_{2,t+1} &= N_{2,t}e^{-\left(F_{2,t}+M_{2,t}\right)}(1-T_{21,t}) + N_{1,t}e^{-\left(F_{1,t}+M_{1,t}\right)}T_{12,t}
\end{align*}
\]

\[C_{1,t} = \frac{F_{1,t}}{F_{1,t}+M_{1,t}}N_{1,t}(1-e^{-\left(F_{1,t}+M_{1,t}\right)})
\]

\[C_{2,t} = \frac{F_{2,t}}{F_{2,t}+M_{2,t}}N_{2,t}(1-e^{-\left(F_{2,t}+M_{2,t}\right)})
\]

with \(N_{1,t}\) and \(N_{2,t}\), the fish numbers in zones 1 and 2 during time step \(t\); \(F_{1,t}\) and \(F_{2,t}\), the fishing mortality coefficients in zones 1 and 2 corresponding to numbers of fish and efforts in each zone during the time step \(t\); \(M_{1,t}\) and \(M_{2,t}\), the natural mortality coefficients in zones 1 and 2 during the time step \(t\); \(T_{12,t}\), the migration rate from zone 1 to zone 2 during time step \(t\) and \(T_{21,t}\), the migration rate from zone 2 to zone 1 during time step \(t\).

No hypothesis is made about the distance between zones or their connexity. The model is not spatially explicit, it is based on the fundamental assumption of homogeneity within each zone (each zone is supposed to behave as a whole).

The case of a single cohort only is considered in the following of the paper. In the case of a single cohort, the time \(t\) is redundant with the age.

System (1) has two unknown state variables \(N_{1,t}\) and \(N_{2,t}\), two measurable output variables \(C_{1,t}\) and \(C_{2,t}\) and six parameters \((M_{1,t}, M_{2,t}, F_{1,t}, F_{2,t}, T_{12,t}, T_{21,t})\). Without external information, natural mortality coefficients \(M_{1,t}\) and \(M_{2,t}\) can be arbitrarily set as it is normally done in cohort analysis. The fishing mortality applied to the whole stock \(F_{i}\) is estimated by VPA and is used to determine the fishing mortality in each zone \(F_{1,t}\) and \(F_{2,t}\). By using the catchability per surface unit \(q'\)

(Laurec and Le Guen, 1981) and the assumption of spatial homogeneity of fish density in each zone \(i\), we can write:

\[q' = \frac{N_{i,t}S_{i,t}}{f_{i,t}}
\]

with \(S_{i,t}\) the surface covered by the stock in zone \(i\) during time step \(t\) (see 1.3.). Because \(q'\) is constant in space, we can write:

\[F_{i,t}S_{i,t} = q' f_{i,t} S_{i,t}
\]

then:

\[F_{i,t} = \frac{S_{i,t}}{f_{i,t}} F_{i,t}
\]

(\(q'\) the catchability per surface unit \(q'\), is defined as being equal to the local CPUE \(U_{a}\), divided by the local density \(D_{a}\), both calculated on a unit surface \(ds\) during time \(t\):

\[q' = \frac{U_{a}}{D_{a}} = \frac{U_{a}}{N_{a}ds} = q_{a}ds
\]

with \(N_{a}\) and \(q_{a}\), the fish number and the catchability in the unit surface \(ds\) during time \(t\).

Contrary to the catchability \(q\), the assumption that the coefficient \(q'\) is spatially constant at a given time \(t\) can be done. Then, catch per unit effort in each zone \(i\) at time \(t\) can be expressed as follow:

\[\text{CPUE}_{i,t} = \frac{C_{i,t}}{f_{i,t}} = \frac{q' \sum_{i} D_{i,t} f_{i,t} ds}{\sum_{i} f_{i,t} ds}
\]

with \(f_{i,t}\) the effort in the elementary surface \(ds\) at time \(t\).

With the assumption of spatial homogeneity of fish density in each zone \(i\) \((D_{a}=D_{a})\), the previous equation can be simplified:

\[\text{CPUE}_{i,t} = q' D_{a} = q' \frac{N_{i,t}}{S_{i,t}} = q_{a} N_{i,t}
\]

with \(N_{i,t}\), the number of fish in zone \(i\) at time \(t\). Such a simplification could be made at the global stock level only under the generally false hypothesis of equality of fish densities in all zones at a given time \(t\):

\[\text{CPUE}_{i} = \frac{C_{i}}{f_{i}} = \frac{q' \sum_{i} (D_{i,t} f_{i,t})}{\sum_{i} f_{i,t}}
\]
Using the conservativity of the number of fish \((N_{t+1,i}+N_{t+1,j}-N_{t+1,i}=0)\), system (1) gives:

\[
g(F_{1,i}) = \frac{C_{l,i}(M_{1,i} + F_{1,i})}{F_{1,i},(1-e^{-F_{1,i}+M_{2,i}})} + \frac{C_{l,i}(M_{2,i} + F_{2,i})}{F_{2,i},(1-e^{-F_{2,i}+M_{2,j}})} - \frac{C_{l}(M_{i} + F_{i})}{F_{i},(1-e^{-F_{i}+M_{j}})} = 0
\]

with \(F_{2,i}\) given by equation (3), \(M_{1,i}\), \(M_{2,i}\), \(C_{l,i}\), \(C_{2,j}\) and \(C_{l}\) already known and \(F_{i}\) estimated by VPA.

Equation (4) is solved numerically with the hybrid method proposed by Press et al. (1994). When \(F_{1,i}\) is assessed, equation (3) is used to calculate \(F_{2,i}\). Next, fish numbers \(N_{1,i}\) and \(N_{2,i}\) are calculated using the catch equations in each zone of the system (1).

The only remaining unknown parameters are \(T_{12,t}\) and \(T_{21,t}\). But system (1) is still not identifiable, as far as infinity of couples \((T_{12,t}, T_{21,t}\) can be a solution. It is under-determined and must be simplified \((i.e.: a\ \text{relation}\ \text{between} \ T_{12,t}\ \text{and} \ T_{21,t}\ \text{has} \ \text{to} \ \text{be}\ \text{found})\). For this, fish movement is supposed to have two components: a random one \(D_{s}\) (diffusion) which is spatially isotropic (brownian motion) and depends only on age and time \(t\), and a deterministic one \(T_{l}\) (advection) which is spatially directed (Okubo, 1980; Deriso et al., 1991; Kleiber and Fonteneau, 1994).

With such assumptions, transfer coefficients \(T_{12,t}\) and \(T_{21,t}\) can be expressed as follows:

\[
T_{12,t} = D_{s} + T'_{12,t} \quad \text{or} \quad T_{21,t} = D_{s} + T'_{21,t} \\
T_{21,t} = D_{s} + T'_{21,t}
\]

\(T_{21,t}\) and \(T_{21,t}\) can not be simultaneously positive because advection is considered as an homogeneous and univocal phenomenon. Now, the unknown parameters are: \((D_{s}, T'_{12,t})\) or \((D_{s}, T'_{21,t})\). In the present study, the coefficient \(D_{s}\) is supposed to be estimated with an auxiliary method (Kleiber et al., 1994; Porch, 1996), neglected compared to \(T'_{12,t}\) or bounded in an interval. Migration rates at terminal age \(T_{12,t}\) and \(T_{21,t}\) are supposed to be equal to values at the previous time step: \(T_{12,t-1}\) and \(T_{21,t-1}\). For all the other time steps, equations (1) and (5) are used to estimate \(T_{12,t}\) and \(T_{21,t}\):

\[
T_{12,t} = \frac{N_{2,t+1,i} + N_{2,t+1,j} - e^{(F_{1,i}+M_{2,i})}}{N_{1,i} - e^{(F_{1,i}+M_{2,i})}}(D-1) \\
T_{21,t} = D_{s}
\]

If \(T_{12,t} \leq D_{s}\ (\Leftrightarrow T'_{12,t} \leq 0\ \text{which} \ \text{is} \ \text{impossible} \ \text{by} \ \text{definition})\), symmetrical equations are used to estimate \(T_{12,t}\) et \(T_{21,t}\).

### 1.2. Extending the method to several fleets

For \(i = 1, \ldots, n\) fleet, fishing mortality on the whole stock can be split up into fishing mortality for each fleet:

\[F_{i} = \frac{C_{l,i}}{C_{l}}, F_{i}\]

So, equation (3) can be used for each fleet:

\[F_{i,2,t} = \frac{S_{i,2,t}, F_{i,2,t}}{S_{2,t}, F_{i,2,t}}, F_{i,3,t}\]

with \(F_{i,3,t}\) the fishing mortality exerted by fleet \(i\) in zone \(j\) (on the local stock \(N_{i,j}\)) at time \(t\).

And equations (3) is changed in \(n\) equations systems (6):

\[
F_{1,2,t} = \frac{S_{1,t}, F_{1,2,t}}{S_{2,t}, F_{1,2,t}}, F_{1,1,t} \\
\vdots \\
F_{n,2,t} = \frac{S_{n,t}, F_{n,2,t}}{S_{2,t}, F_{n,2,t}}, F_{n,1,t}
\]

then,

\[F_{2,t} = \sum_{i} F_{2,t,i} = \sum_{i} S_{i,t}, F_{2,t,i}, F_{i,1,t} = \sum_{i} \left( \frac{S_{i,t}, F_{2,t,i} - C_{i,1,t}}{S_{2,t}, F_{i,1,t} - C_{i,1,t}} \right), F_{i,1,t}\]

Equation (4) and (7) are fully determined and can be numerically solved.

### 1.3. Application to Atlantic yellowfin tuna

To illustrate the method, it is applied to yellowfin tuna of Atlantic considering the eastern and western fraction \(s\) of the stock. At the time of the present work the most
recent fisheries statistics was not reliable and the present run must be considered as a first one, which purpose is just to check the method. More analysis are actually performed and will be further presented.

ICCAT catches and effort data are used. Age decomposition is performed by month for the for first age groups (0-1-2-3) using the age-length adjustment method developed by Gascuel (1994) with a two stage growth model (Gascuel et al., 1992). Because the yellowfin mean size at age is close for older age groups, slicing is used to distinguish the two last age groups (4-5+) with ICCAT limit. Catches and effort data are averaged on the period 198***-199 for the eastern fleet and 198-199 for the western fleet. A backward VPA is then runned with a quarterly time step as it is usually performed by ICCAT working groups.

The spatial VPA methodology also needs an estimation of stock surface in each zone as an auxiliary information (equation (3)). Stock surface is a theoretical concept as far as a stock has no clear delimited geographic frontiers. To avoid variations with total abundance, it can be defined as the surface of a given fraction of the stock (Swain and Sinclair, 1993) but it is most of the time considered as the area of the stock where density is higher than a given threshold (Swain and Wade, 1992, Marshall and Frank, 1994). In the present study, long line catches were used to estimate yellowfin stock surface in both east and west ICCAT zones. Catches were cumulated by month and 5°x5° square over the whole 1956-1993 period. The total catches were spatially interpolated to produce a continuous map and two catches level (150 and 300 tons by 5°.month) were arbitrarily defined as the stock limit. Such levels were used to calculate yellowfin stock surface in each side of the Atlantic ocean (fig. 1). This method was used because the surface of a 5°x5° square varies with latitude and then the number of 5°x5° square can’t be used. All calculations were computed with a GIS (savane software ©orstom, 1995).

2. RESULTS

2.1. Global and local fishing mortalities

2.2. Fish numbers in each zone and migration rates
3. DISCUSSION AND CONCLUSION

*** A monthly basis is certainly more accurate for the VPA performed at the global stock level (Fonteneau, 1994b) and a short time step is necessary for the spatial VPA because of the important time variability of migration phenomena.***

Spatial VPA results emphasize the hypothesis of massive migrations across the Atlantic ocean. Mark recapture data (Bard) and length frequency analyses (Fonteneau) already suggested such an hypothesis. Foucher (199) give a first rough approximation of exchange rates but the method he uses is only based on purse seiners catches and doesn’t take into account stock surface. Our results suggest that yellowfin recruitment is on average not equally distributed between east and west Atlantic. % of the recruitment comes from in eastern zone, very likely from the main reproduction area of the gulf of Guinea (Fonteneau, 1993) and probably from accessory breeding areas such as Cape Verde islands (Santa Rita Vieira, 199). % of the recruitment could come from western zones, ............

The relationship between commercial vessels CPUE and stock abundance is generally not a simple linear function. Many well known phenomena lead to a non linear relationship between CPUE and fish abundance. They are generally attributed
to the resource spatial heterogeneity and to the fishermen search behaviour (Clark et al., 1979 ; Hilborn and Walters, 1987 ; Hilborn and Walters, 1992 ; Gauthiez, 1997). The consequences of such phenomena on the spatial heterogeneity of catchability at age $q_t$ (and more precisely on the variability of the catchability by surface unit $q_{t}'$, at a given time $t$ between different given zones) are less studied.

The catchability $q_{t}'$ generally varies with stock abundance (its variations are density-dependent) and with fishermen behaviour (Hilborn and Walters, 1992 ; Gauthiez, 1997). Because fish density is spatially heterogeneous, catchability $q_{t}'$, also varies in space. These aspects are not detailed here, but it is important to keep in mind the potential consequences of violation of the spatial catchability $q_{t}'$, homogeneity hypothesis.

Nevertheless, simulations show that in many cases the spatial VPA method is reliable (Maury et al., 1997 a and b). It seems to be biased when simultaneously fishing effort is highly deterministic and fish density is highly heterogeneous and density dependent. Different statistical tools such as mean/variace relationship (Gauthiez, 1997) or geostatistical selectivity curves (Petitgas, 1994) may allow to identify such « dangerous » cases. In those cases, the use of different stock surface estimates allows to bound spatial VPA outputs between a high and a low value (Maury et al. 1997 a and b).

**CONCLUSION**

The method and its underlying hypothesis are tested for two zones on several data sets simulated with the advection-diffusion based simulator SHADYS (Simulateur HALieutique de DYnamiques Spatiales) (Maury and Gascuel, 1997).

The proposed method enables to estimate local parameters which are usually estimated at the global stock level (fish numbers and fishing mortality coefficients).

Le découpage n’est pas biologiquement pertinent, il ne l’est que pr estimator des taux d’interaction entre pêcheries est et ouest , le schema de migration non plus. Il faut changer de modele

**RÉFÉRENCES**


Marshall and Frank, 1994


Swain and Sinclair, 1993

Swain and Wade, 1992